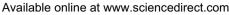


ANIMAL BEHAVIOUR, 2008, **76**, 87–95 doi:10.1016/j.anbehav.2008.01.012







The genetic similarity between pair members influences the frequency of extrapair paternity in alpine marmots

AURÉLIE COHAS*†, NIGEL GILLES YOCCOZ‡, CHRISTOPHE BONENFANT*, BENOÎT GOOSSENS§, CÉLINE GENTON*, MAXIME GALAN**, BART KEMPENAERS† & DOMINIQUE ALLAINÉ*

*Laboratoire Biométrie et Biologie Evolutive, UMR CNRS 5558, Université Claude Bernard Lyon 1 †Department of Behavioural Ecology & Evolutionary Genetics, Max Planck Institute for Ornithology,

Seewiesen

‡Department of Biology, University of Tromsø §Biodiversity and Ecological Processes Group, Cardiff School of Biosciences

**Centre de Biologie et de Gestion des Populations, UMR INRA - IRD - Cirad, Montpellier

(Received 2 February 2007; initial acceptance 9 March 2007; final acceptance 9 January 2008; published online 12 May 2008; MS. number: 9260R)

Extrapair paternity is widespread in birds and mammals. In particular, the alpine marmot, *Marmota marmota*, has a high frequency of extrapair paternity that seems to be explained by the genetic compatibility hypothesis. We investigated whether the number and proportion of extrapair young depend on the heterozygosity (individual genetic diversity) of the social male, or on the genetic similarity between the social male and his mate (relatedness). Both the number and the proportion of extrapair young increased with both high similarity and dissimilarity between the social pair. In combination with previous results, our study suggests that patterns of extrapair paternity in alpine marmots can best be explained by the genetic compatibility hypothesis, and more precisely its optimal outbreeding variant. Our results indeed suggest that extrapair paternity is a mechanism to avoid both in- and outbreeding depression. We discuss which proximal mechanisms may be involved in extrapair paternity in this species.

© 2008 The Association for the Study of Animal Behaviour. Published by Elsevier Ltd. All rights reserved.

Keywords: alpine marmot; cooperative breeding; inbreeding; *Marmota marmota*; mating system; miscrosatellite; related-ness; reproductive skew

87

The genetic compatibility hypothesis states that females may benefit from extrapair paternity (EPP) if their genes are more compatible with those of the extrapair male than those of their pair mate (Zeh & Zeh 1996; Tregenza &

Correspondence: A. Cohas, Laboratoire Biométrie et Biologie Evolutive, UMR CNRS 5558, Université de Lyon, Université Claude Bernard Lyon 1, 43 Bd du 11 novembre 1918, 69622 Villeurbanne cedex, France (email: cohas@biomserv.univ-lyon1.fr). N. G. Yoccoz is at the Department of Biology, University of Tromsø, N-9037 Tromsø, Norway. B. Goossens is at the Biodiversity and Ecological Processes Group, Cardiff School of Biosciences, Cardiff University, PO Box 915 Cathays Park, Cardiff CF10 3TL, U.K. M. Galan is at the Centre de Biologie et de Gestion des Populations, UMR INRA – IRD – Cirad – Montpellier SupAgro, Campus international de Baillarguet, CS 30016, 34988 Montferrier-sur-Lez cedex, France. B. Kempenaers is at the Max Planck Institute for Ornithology, Postfach 1564, 82305 Starnberg/Seewiesen, Gemany. Wedell 2000). Inbreeding is one form of genetic incompatibility, so females may benefit from EPP by decreasing the inbreeding level of their offspring (Zeh & Zeh 1996; Tregenza & Wedell 2000). Although still poorly supported (Kempenaers 2007), this inbreeding avoidance hypothesis has received some support in both socially monogamous birds (Blomqvist et al. 2002; Foerster et al. 2003) and mammals (Sillero-Zubiri et al. 1996).

The genetic benefits that females actually gain from their extrapair mate choice depend on their ability to bias egg fertilization in favour of the best male. Theoretically, this could be achieved via pre- and postcopulatory choice mechanisms (Pizzari & Birkhead 2002). The number and proportion of extrapair young (EPY) should then increase as the genetic quality or the genetic compatibility of the social mate decreases relative to that of other available sexual partners. If females are not able to adjust fertilization relative to male quality/compatibility, or if females seek extrapair copulations (EPCs) to increase their offspring's genetic diversity, this relation is no longer expected. Thus, to understand better the evolution of EPP, it is crucial to examine the distribution of EPY within and among litters in relation to the characteristics of both social and extrapair males.

The alpine marmot, Marmota marmota, is a socially monogamous, territorial species where only dominant individuals are thought, from behavioural observation, to obtain access to reproduction (King & Allainé 2002; Hackländer et al. 2003) and EPP occurs frequently (Goossens et al. 1998; Cohas et al. 2006). Female choice appears to be an important determinant of EPP in alpine marmots as indicated by the nonrandom distribution of EPP across males (within-pair males and extrapair males, Cohas et al. 2006). Therefore, the alpine marmot is a suitable species for investigating the underlying ultimate causes that drive EPP and previous results seem to support the genetic compatibility hypothesis. Indeed, (1) the occurrence of EPP among litters depends on the genetic similarity between social mates (Cohas et al. 2006), (2) extrapair mates are more heterozygous and less genetically similar to the female than the corresponding within-pair mate (social partner; Cohas et al. 2006, 2007a), and (3) EPY survive better and have better access to reproduction than withinpair young (WPY) (Cohas et al. 2007b).

In alpine marmots, dispersal is costly because the probability of surviving and acquiring a breeding vacancy decreases rapidly with dispersal distance (Frey-Roos 1998). Consequently, 12–22% of subordinates become dominant in their natal territory and about 50% of subordinates become dominant in the immediate neighbourhood (Frey-Roos 1998; Magnolon 1999). This dispersal pattern leads to higher relatedness among pair mates than among random pairs (Cohas et al. 2006). Moreover, homozygous juveniles survive less well than heterozygous ones, especially under harsh winter conditions (Da Silva et al. 2006). Thus, the conditions for inbreeding avoidance to be a strong constraint driving EPP are met in the alpine marmot.

We focused on the inbreeding avoidance variant of the genetic compatibility hypothesis, specifically on the distribution of EPY among and within litters in relation to the male's genetic characteristics. In particular, we investigated whether the presence, the number and the proportion of EPY depended on within-pair mate heterozygosity (as an indicator of individual quality) and within-pair mate genetic similarity to the female.

METHODS

Study Species and General Procedures

Alpine marmots are territorial cooperative breeders with reproduction highly skewed towards dominant individuals (King & Allainé 2002; Hackländer et al. 2003). The basic social unit is a family group of 2–20 individuals with a dominant breeding pair, sexually mature subordinates (at least 2 years old), yearlings and juveniles (Perrin et al. 1993). The physiological reproductive functions of the great majority of sexually mature subordinate females (Hackländer et al. 2003) and of sexually mature subordinate males (Arnold 1990; Goossens et al. 1998; Cohas et al. 2006) are inhibited by aggressive behaviour by the dominant individual of the same sex. Hence most subordinates disperse from 2 years old onwards and become transient individuals in search of a breeding territory (Frey-Roos 1998; Magnolon 1999).

The study site is located in La Grande Sassière Nature Reserve (French Alps, 45°29'N, 6°59'E, 2300 m above sea level). It covers 40 ha of alpine open meadows. From 1990 to 2006, we captured marmots from early April to late July on at least 45 days a year. We used two-door, live-capture traps baited with dandelion, Taraxacum densleonis. We placed the traps near the entrance of the main burrows of each family group to assign trapped individuals to their family and checked them every half hour to limit the time an individual was trapped. Once captured, individuals were tranquillized with Zolétil 100 (0.1 ml/kg) and individually marked with a numbered eartag $(1 \text{ cm} \times 3 \text{ mm})$ and a transponder (model ID100, 0.9 cm long, <0.1 cm in diameter, Trovan Ltd, www.Trovan.com) injected under the skin of the neck for permanent individual recognition. In addition, a piece of coloured plastic (<1 cm²) was fixed to one ear. Trapped individuals were sexed and aged by size (up to 3 years of age). Morphological characteristics (presence of visible testis for males, developed teats for females, development of scent glands for both sexes) were used to confirm individual social status determined from observation. For genetic analyses, we collected hair from all individuals captured since 1992, and tissue biopsies from the flank of individuals since 1997; a piece of skin (<1 mm³) was removed with a biopsy punch (Alcyon, Lyon, France). The marking and biopsies did not cause any bleeding. Handling lasted a maximum of 10 min and individuals were absent from their territory for a maximum of 30 min. We never observed exclusion from the territory for any individual of any age following capture.

We determined the composition of 20 families from capture histories combined with intensive observations using 10×50 binoculars and 20×60 telescopes from a distance of 80-200 m. Each family was observed on average 1 h per day for a minimum of 30 h per year with 1 h sessions being randomly distributed during the period of activity from 0800 to 1200 hours and from 1500 to 2100 hours. We recorded the number of yearlings, 2-year-olds and adults of each sex and their social status for each family. Size allowed us to categorize individuals as yearlings, 2-year-olds or adults, and scentmarking behaviour and aggressive interactions allowed us to categorize individuals as subordinates or dominants (Bel et al. 1999). From additional daily observations, we recorded the date and the litter size at emergence $(X\pm SD=3.6\pm 1.2, \text{ range } 1-7)$. Virtually all emerged juveniles were trapped within 3 days of emergence (Allainé et al. 2000; Allainé & Theuriau 2004), and none were rejected by their mother or family group after capture. All manipulations were approved by the Centre National de la Recherche Scientifique and the Vanoise National Park Authority.

Genotyping and microsatellite characteristics

In total, 797 individuals were typed at 16 microsatellite loci: SSBibl1, SS-Bibl18, SS-Bibl20, SS-Bibl31, SS-Bibl4 (Klinkicht 1993); MS41, MS45, MS47, MS53, MS56, MS6, ST10 (Hanslik & Kruckenhauser 2000); Ma002, Ma018, Ma066, Ma091 (Da Silva et al. 2003). Details of the genotyping method are given in the Appendix. Depending on the locus, 97.1–99.8% of all individuals were genotyped (Table 1). To estimate the genotyping error rate, we typed 96 randomly chosen individuals for each microsatellite locus. No discrepancy between the two genotypes was found, so the probability of finding an error for one allele should not exceed 0.0003.

Table 1 summarizes the microsatellite characteristics. Using the library 'adegenet' for R (Jombart 2007), we carried out Hardy–Weinberg tests on dominant adults only, to avoid potential bias caused by family structure, and on all cohorts pooled to ensure sufficient sample size (N = 160). Except for Ma002 ($\chi^2 = 308.90$, 10000 replicates, P = 0.027), none of the loci showed deviation from Hardy–Weinberg equilibrium (1 > P > 0.064).

Parentage analysis

The genotypes of each offspring and of the supposed parents were compared to check maternity. For 16×645 mother-offspring comparisons, only one mismatch at one locus (SS-Bibl20) between the supposed mother (usually the dominant female) and one of its offspring was found. We then defined young as WPY if no mismatch was observed with the dominant male genotype (553 WPY of 595 offspring). We defined young as EPY if at least one mismatch was observed with the dominant male genotype (42 EPY: 1-8 mismatches). For 13 offspring, exclusions of paternity were based on only one mismatch with the social male. We consider it unlikely that these young were WPY, because (1) the genotyping error rate was low (0.0003), (2) all these offspring and their parents were retyped and their genotypes confirmed, (3) the average mutation rate for microsatellites is low $(1.67 \times 10^{-4} \text{ per})$ generation in *M. marmota*, Rassmann et al. 1994) and (4) only a single mismatch with the putative mother was found (see above).

The genotypes of EPY were then compared to the genotypes of all known sexually mature males and two types of EPY were defined. Within-group EPY had geno-types compatible with the genotype of a subordinate male in their family but incompatible with the genotypes of all males from outside their family (13 within-group EPY). Extragroup EPY had a genotype incompatible with those of all subordinate males of their family. The biological father was known for 13 extragroup EPY and unknown for 16.

Paternity analysis was repeated using the Cervus 3.0.3 software (Kalinowski et al. 2007), with the following settings: 20 candidate parents of each sex per offspring; 95% of candidate parents sampled; error rate of 1% to allow for mistyping and for mutations or null alleles (Table 2); and assignment at a 95% confidence level. For

	Bibl1	Bibl18	8	Bibl20	Bibl31	Bibl4	Ma002	Ma018	Ma066	Ma091	MS41	MS45	MS47	MS53	MS56	MS6	ST10	
•	Alleles Freq	Alleles	Freq Alk	eles Freq	Alleles Frec	Alleles Freq Alleles Freq Alleles Freq	Alleles Freq		Alleles Freq Alleles Freq Alleles Freq		Alleles Freq	Alleles Freq Alleles Freq Alleles	Alleles Freq A	Freq	Alleles Freq	Alleles Freq Alleles	Vleles Freq	Total
~~~~~~	95 0.16 97 0.20 101 0.45 103 0.45 103 0.14 107 0.14 109 0.05	133 137 145 145 147	<pre>&lt;0.01 206 &lt;0.01 208 0.01 208 0.03 216 0.13 218 0.39 220 0.31 222 0.11 222</pre>	6 0.01 8 0.22 6 0.38 8 0.31 0 0.08 2 <0.01	157 0.47 159 0.30 161 0.18 163 0.05	7 175 0.12 0 178 0.00 8 188 0.18 5 190 0.67 192 0.03	265 <0.01 271 0.21 273 <0.01 273 <0.01 279 0.46 281 0.32 283 <0.01	296 0.27 298 0.73	231 0.62 233 0.03 241 0.35	159 0.13 167 0.09 171 0.05 171 0.02 173 0.15 175 0.43 177 0.02 179 0.09 188 <0.01	184 0.17 186 0.83	107 0.36 109 0.53 111 0.11	176 0.04 180 0.23 182 0.19 184 0.17 186 0.34 188 0.33 190 0.01	132 0.13 140 0.47 142 0.39 144 <0.01 148 <0.01	104 0.02 106 0.26 108 0.72 110 <0.01	142 0.06 158 0.87 160 0.07	116 0.14 118 0.29 120 0.19 130 0.05 132 0.15 134 0.15 136 0.04	
z	795	792	2	786	795	793	774	777	786	789	783	792	787	789	787	785	785	
NO. OT alleles	9	9		9	4	5	9	2	m	6	2	£	7	5	4	£	7	4.88
PIC	0.67	0.63	3	0.65	0.59	0.46	0.57	0.32	0.39	0.72	0.24	0.49	0.73	0.53	0.35	0.22	0.79	0.523
NE-1P	0.70	0.73	3	0.72	0.77	0.87	0.79	0.92	0.88	0.63	0.96	0.83	0.63	0.81	0.91	0.97	0.54	0.020
NE-2P	0.52	0.57	7	0.55	0.61	0.72	0.65	0.84	0.79	0.44	0.88	0.71	0.45	0.68	0.81	0.88	0.37	0.0007
NE-PP	0.34	0.39	6	0.37	0.44	0.56	0.50	0.75	0.68	0.24	0.80	0.57	0.27	0.54	0.70	0.78	0.19	< 0.0001

**Table 2.** Assignment rate and reliability of assignment (%) at a 95% level of confidence given the set of 16 microsatellites used in the parentage analysis and assuming 20 candidate parents for each sex, 95% of candidate parents sampled and an error rate of 1%

	Father, given known mother		Parent pair with sex known
Assignment rate Identity of most likely candidate	91.0	95.0	93.0
True parent(s) Nonfather/ mother/parents (true parent(s) sampled)	98.3 0.005	98.6 0.8	95.0 0.7
Nonfather/ mother/parents (true parent(s) unsampled)	0.01	0.6	2.1*, 2.2†, 0.02‡

*True mother unsampled.

†True father unsampled.

‡Neither true parent sampled.

27 young from 20 litters, some of the WPY could also be assigned to both the dominant and a subordinate male. However, for 14 young the dominant male had a higher likelihood of being the father than the subordinate male. Moreover, in all cases, the sexual organs of the subordinate male showed no sign of development at capture, and the rest of the litter could always be assigned to the dominant male. Thus, we parsimoniously considered all these young as WPY. There were no ambiguities in the assignments of EPY.

#### Estimates of heterozygosity and genetic similarity

Individual genetic diversity was estimated as the standardized individual heterozygosity (SH), which is defined as the proportion of heterozygous loci divided by the mean heterozygosity of the scored loci (Coltman et al. 1999). This method is used to account for the fact that few individuals were scored at fewer than 16 loci.

Three estimators of genetic similarity between mates were calculated using functions written in R (R Development Core Team 2007): Queller & Goodnight's estimator ( $R_{QG}$ , Queller & Goodnight 1989), Lynch & Ritland's estimator ( $R_{LR}$ , Lynch & Ritland 1999) and Identity (I, Belkhir et al. 2002). The three estimators were highly correlated ( $R_{QG}$  versus I: r = 0.75,  $R_{QG}$  versus  $R_{LR}: r = 0.86$ ,  $R_{LR}$  versus I: r = 0.86, N = 316410 pairs) and the results found were independent of the estimator used. Therefore, we present only the results using Queller & Goodnight's estimator.

#### **Definitions and Data Analyses**

For this study, we limited the analyses to 103 litters where dominant individuals and all offspring were known and genotyped (N = 369), and where the number of mature male subordinates present in the family group was known. Litters composed of only WPY were defined as within-pair litters, those composed exclusively of EPY

were defined as extrapair litters, and those containing both WPY and EPY are referred to as mixed litters.

We investigated the effect of fixed terms (number of mature male subordinates present in the family group, male heterozygosity, genetic similarity between social partners, litter size) on the presence, number and proportion of EPY within litters (three related variables). To account for repeated measures (26 pairs were present in 2-5 years), we used generalized estimating equations (GEE) to estimate the parameter of generalized mixed models (Liang & Zeger 1986; Zeger & Liang 1986; Diggle et al. 2002). This procedure is likely to be more robust than likelihood-based estimation for generalized mixed models because it makes broad hypotheses about data structure, and is better adapted to departure from normality of random effects and small sample sizes within clusters (Carlin et al. 2001). We used GEE with the pair as the clustering factor. We chose an exchangeable correlation matrix to specify the same correlation between all observations of the same cluster (this is analogous to the correlation structure derived from assuming a random factor in a mixed model; Horton & Lipsitz 1999). Moreover, we included the number of sexually mature subordinate males present in the family group as a fixed variable in all models since it affects the occurrence of EPP in alpine marmots through male-male competition (Cohas et al. 2006).

In the model with the number of EPY as the dependent variable, we used GEE with a logarithm link and a variance given by a gamma distribution to account for overdispersion (see Results, Venables & Ripley 2002). In the other two models with the presence or the proportion of EPY as the dependent variable, we also used GEE but with a logit link and a variance given by a binomial distribution (i.e. same link and variance function as in a logistic regression). The significance of fixed terms was assessed using the robust *z* statistics of parameter estimates (Diggle et al. 2002).

For all statistical analyses we used R 2.5.1 software and the gee library (R Development Core Team 2007). Unless otherwise stated, all tests were two tailed, the level of significance was set to 0.05, and parameter estimates are given  $\pm$ SE.

## RESULTS

#### Frequency of Extrapair Paternity

For litters where dominant individuals and all offspring were known and genotyped and where the number of mature male subordinates present in the family group was known, 36 (9.7  $\pm$  1.5%) of 369 offspring were EPY and 20 (19.4  $\pm$  3.9%) of 103 litters contained at least one EPY (Table 3). Ten (50.0  $\pm$  11.1%), six (30.0  $\pm$  10.2%), three (15.0  $\pm$  8.0%) and one (5.0  $\pm$  4.9%) of these litters contained one, two, three and five EPY, respectively, the mean number of EPY per litter was 1.8  $\pm$  0.2 and half of these litters containing at least one EPY contained more than 50% of EPY.

EPY were not randomly distributed (i.e. did not follow a Poisson distribution) both among all litters

			EPY		
	WPY	- Within-group*	Extragroup†	Unidentified‡	Total
Within-pair litters	298 (83)	_	_	_	298 (83)
Mixed litters	35 (15)	4 (2)	8 (4)	10 (9)	57 (15)
Extrapair litters		9 (4)	5 (1)	0 (0)	14 (5)
Total	333 (98)	13 (6)	13 (5)	10 (9)	369 (103

Table 3. Distribution of within- and extrapair offspring among the different types of litters

Numbers of litters are given in parentheses. WPY: within-pair young; EPY: extrapair young.

*Young sired by a subdominant male within the social group.

†Young sired by a known male from outside the social group.

[‡]Young sired by an unidentified extrapair male outside the social group.

(Kolmogorov–Smirnov test: D = 1.89, N = 103, P < 0.001) and among litters containing at least one EPY (D = 3.55, N = 20, P < 0.001).

## **Factors Influencing Loss of Paternity**

The presence of EPY within a litter increased with the number of sexually mature subordinate males present in the family group (Table 4, Fig. 1). The presence of EPY in a litter also depended on a quadratic relation with the genetic similarity between pair mates. Thus, the probability of observing EPY in a litter was higher for extreme (low and high) values of genetic similarity between mates than for intermediate ones (Table 4, Fig. 1). Similarly, both the number and the proportion of EPY per litter followed a quadratic relation with the genetic similarity

**Table 4.** Generalized estimating equation models showing the effects of terms on the presence of extrapair young (EPY) in a litter, the number of EPY per litter and the proportion of EPY per litter

	Coefficients	SE	Р
Presence of EPY in litter			
Number of subordinate males	0.980	0.284	<0.001
Within-pair male heterozygosity	1.708	1.565	0.275
Genetic similarity between pair mates		1.130	0.300
(Genetic similarity between pair mates) ²	6.709	2.654	0.011
Litter size	0.191	0.185	0.302
Number of EPY per litter			
Number of subordinate males	0.285	0.167	0.087
Within-pair male heterozygosity	1.176	0.962	0.221
Genetic similarity between pair mates	-1.333	0.548	0.015
(Genetic similarity between pair mates) ²	3.603	1.516	0.017
Litter size	0.188	0.164	0.252
Proportion of EPY per litter			
Number of subordinate males	0.383	0.209	0.067
Within-pair male heterozygosity	1.278	2.622	0.238
Genetic similarity between pair mates	-1.637	0.893	0.067
(Genetic similarity between pair mates) ²	5.181	2.279	0.023
Litter size	-0.107	0.217	0.620

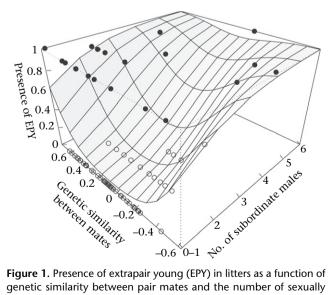
between pair mates: EPY were more numerous for extreme (low and high) values of genetic similarity between mates than for intermediate values (Table 4, Fig. 2).

The presence of EPY within a litter and the number and proportion of EPY were independent of within-pair male heterozygosity and of litter size (Table 4). The number and proportion of EPY within a litter were also independent of the presence of sexually mature subordinate males in the family group (Table 4).

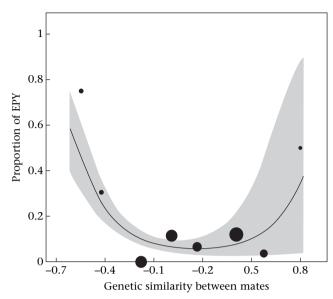
## Patterns of Paternity Gain

Of the 20 litters containing at least one EPY, two were fathered by three males, whereas all others were fathered by two males. The within-pair male sired none of the young in only five  $(4.5 \pm 2.1\%)$  litters (Table 3). Among the five within-pair males with complete paternity loss only one never sired young in other years.

Thirteen of 36 EPY ( $36.1 \pm 8.0\%$ ), born in three ( $15.0 \pm 8.0\%$ ) of the 20 litters containing at least one EPY, were fathered by an extrapair male within the same family. The remaining 23 ( $63.9 \pm 8.0\%$ ) EPY, born in 17



**Figure 1.** Presence of extrapair young (EPY) in litters as a function of genetic similarity between pair mates and the number of sexually mature subordinate males present in the family group. ●: Litters with EPY; ○: litters without EPY; the surface represents the fitted model.

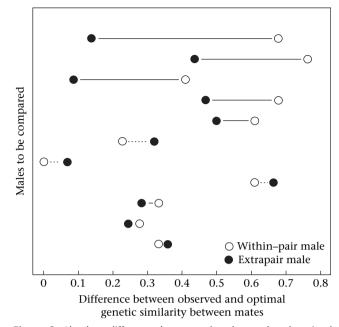


**Figure 2.** Proportion of extrapair young (EPY) per litter as a function of the genetic similarity between pair mates. The black circles represent observed data averaged over classes of genetic similarity between mates (class width 0.2) and their size is proportional to the number of litters within each class. The solid line shows the fitted model and the grey surface represents standard errors of the fitted model.

(85.0  $\pm$  8.0%) of the 20 litters containing at least one EPY, were known or inferred to be fathered by a transient male (Table 3). None of the EPY was fathered by a dominant or a subordinate male from a neighbouring group. Extrapair mates were never closely related to the dominant female, but sometimes they were related to the dominant male: one extrapair male was a brother of the dominant male and sired two young in a litter and five were sons of the dominant male and sired a total of seven young in four litters.

From theory, there are clear expectations with regard to the direction of the difference in heterozygosity and in genetic similarity between extrapair and within-pair males. We thus tested only for extrapair males being more heterozygous and more optimally similar to the females and used one-tailed instead of two-tailed tests. Among the 11 identified extrapair males, eight  $(72 \pm 10.9\%)$  showed higher heterozygosity than that of the corresponding within-pair males: four of six extragroup extrapair males and four of five within-group extrapair males were more heterozygous than the corresponding within-pair males. However, the extrapair males were not significantly more heterozygous than the corresponding with-pair males (one-tailed Wilcoxon signed-ranks test: V = 27, N = 11, P = 0.31).

Seven  $(63 \pm 14.6\%)$  of the 11 identified extrapair males showed a genetic similarity to the female that was closer to the optimum (optimal genetic similarity estimated from the GEE model on EPY proportion = 0.16) than that of the corresponding within-pair mate: three of six extragroup extrapair males and four of five within-group extrapair males were closer to the optimal genetic similarity with the female than the corresponding social males



**Figure 3.** Absolute difference between the observed and optimal genetic similarity between the female and the within-pair and corresponding extrapair males. Optimal genetic similarity equals 0.16, as obtained from the generalized estimating equation model of extrapair young proportion within a litter (see Results).

(Fig. 3). However, the extrapair males only tended to show a more optimal level of genetic similarity to the female than the corresponding social males (one-tailed Wilcoxon signed-ranks test: V = 50, N = 11, P = 0.07).

## DISCUSSION

Previous investigations on the ultimate causes driving EPP in the alpine marmot supported the genetic compatibility hypothesis, and more specifically its inbreeding avoidance variant: the comparison of the genetic characteristics of WPY and EPY half-siblings indicated that extrapair males were less genetically similar to the female than the corresponding within-pair males (Cohas et al. 2007a), and EPY outperformed WPY as indicated by their higher probability of surviving and becoming dominant (Cohas et al. 2007b). Our results matched the prediction of the optimal outbreeding hypothesis. Indeed, the probability that EPY were present and the number and proportion of EPY within a litter increased with high similarity and dissimilarity between the social pair mates. Such a pattern has rarely been reported, and only in bird species such as the pied flycatcher, Ficedula hypoleuca (Rätti et al. 1995) or the house sparrow, Passer domesticus (Bonneaud et al. 2006).

Female choice for extrapair mates does not seem to target heterozygous males ('heterozygosity as good genes', Brown 1997) in the alpine marmot. Indeed, (1) neither the probability that EPY were present nor the number and proportion of EPY within a litter increased with the homo-zygosity of the social mate; (2) extrapair males were not more heterozygous than the corresponding within-pair

males; (3) finally, contrary to a central assumption of this hypothesis (Mitton 1993; Brown 1997), heterozygosity is not heritable in alpine marmots (Cohas et al. 2007a).

Precopulatory mechanisms are important in determining the pattern of EPP in alpine marmots for at least two reasons. First, the fact that the probability that EPY are present depends on the number of sexually mature subordinate males present in the family group suggests that the social setting affects the conditions for EPCs to occur (Cohas et al. 2006). Dominant male marmots (i.e. the within-pair mate) may limit the opportunity of females to find extrapair mates either through reproductive suppression of male competitors or through mate guarding (Arnold & Dittami 1997). Females may thus also be limited by the availability of adequate extrapair males. Second, there is additional evidence that alpine marmots avoid inbreeding before copulation. For example, the level of testosterone of subordinate males decreases with their relatedness to the dominant female but not to that of the dominant male (Magnolon 1999), a pattern also found in prairie voles, Microtus ochrogaster (Carter et al. 1986) and in white-footed mice, Peromyscus leucopus (Wolff 1992). Similarly, the dispersal rate of subordinate females increases with their relatedness to the dominant male but not to that of the dominant female (Magnolon 1999). Moreover, owing to the social structure of the alpine marmot, different potential EPC candidates exist: (1) subordinate males of the family group, (2) subordinate or (3) dominant males of other family groups and (4) transient males in search of a territory. Among these potential extrapair males, we showed that EPP preferentially involved transient individuals, probably originating from distant family groups, and hence less related to the female. Furthermore, the few subordinate males of the family group that obtained EPP were all unrelated to the female. Finally, incestuous matings are extremely rare: in 20 years of study, only one mating involving a mother and her son was observed.

The fact that the number and proportion of EPY within litters depended on the genetic similarity between pair mates can be explained by several mechanisms. First, females may copulate more frequently with an extrapair male or less frequently with their own mate when the latter is more genetically similar (precopulatory process). This assumes that females can assess their genetic similarity with their mate. Alternatively, the frequency of EPP in alpine marmots may also partly be explained by postcopulatory 'choice' mechanisms. For example, if sperm from different males are present within the female's reproductive tract, the sperm of the more compatible male may have an advantage, that is, a higher probability of fertilizing the eggs. More litters with EPY and more identified extrapair males are needed to compare the proportion of young sired in relation to the genetic similarity of the female with both within-pair and extrapair males (e.g. Foerster et al. 2006).

Although the good genes hypothesis has unquestionably received more empirical support than any other hypothesis in birds (reviewed in Westneat et al. 1990; Birkhead & Møller 1992; Griffith et al. 2002), more recently, empirical support for the compatibility hypothesis, and especially the inbreeding avoidance variant of this hypothesis, has started to accumulate (Dobson et al. 1997; Kempenaers et al. 1999; Blomqvist et al. 2002; Foerster et al. 2003; Freeman-Gallant et al. 2003; Eimes et al. 2005; Tarvin et al. 2005). In birds and mammals with group living and cooperative breeding, inbreeding avoidance is a strong evolutionary force (Stacey & Koenig 1990; Jennions & Macdonald 1994; Creel & Macdonald 1995; Hoffman et al. 2007). Indeed, in these species, limited opportunities for dispersal and a scarcity of breeding vacancies result in kin of opposite sex residing in the same group (habitat saturation hypothesis: Selander 1964; Emlen 1982; Brown 1987). Consequently, inbreeding avoidance may be an especially strong evolutionary force driving EPP in group-living species (Malurus splendens: Brooker et al. 1990; Canis simensis: Sillero-Zubiri et al. 1996). Indeed, in mammalian species where paternity analyses have been undertaken, EPP is rarer among species living in solitary pairs than among species living in groups (percentage of litters with EPY: solitary pair median = 2.1; group living median = 20.5; Wilcoxon test: W = 8, N = 14, P = 0.04, Cohas 2006).

#### Acknowledgments

We thank all the students involved in the trapping of alpine marmots at La Grande Sassière, Alexander Girg and Sylvia Kuhn for their help in the genotyping of marmots, G. A. Parker for interesting discussions on postcopulatory mechanisms in marmots and three anonymous referees for constructive comments on the paper. We are grateful to the authorities of the Vanoise National Park for allowing us to work in the Grande Sassière Nature Reserve. Financial support was received from CNRS (France), the Région Rhône-Alpes (XI plan Etat-Région), and the Max Planck Society.

#### References

- Allainé, D. & Theuriau, F. 2004. Is there an optimal number of helpers in alpine marmot family groups? *Behavioral Ecology*, 15, 916–924.
- Allainé, D., Brondex, F., Graziani, L., Coulon, J. & Till Bottraud, I. 2000. Male-biased sex ratio in litters of alpine marmots supports the helper repayment hypothesis. *Behavioral Ecology*, **11**, 507– 514.
- Arnold, W. 1990. The evolution of marmot sociality: II. Costs and benefits of joint hibernation. *Behavioral Ecology and Sociobiology*, 27, 239–246.
- Arnold, W. & Dittami, J. 1997. Reproductive suppression in male alpine marmots. Animal Behaviour, 53, 53–66.
- Bel, M. C., Coulon, J., Sreng, L., Allainé, D., Bagneres, A. G. & Clement, J. L. 1999. Social signals involved in scent-marking behavior by cheek-rubbing in alpine marmots (*Marmota marmota*). *Journal of Chemical Ecology*, 25, 2267–2283.
- Belkhir, K., Castric, V. & Bonhomme, F. 2002. IDENTIX, a software to test for relatedness in a population using permutation methods. *Molecular Ecology Notes*, 2, 611–614.
- Birkhead, T. R. & Møller, A. P. 1992. Sperm Competition in Birds: Evolutionary Causes and Consequences. London: Academic Press.

- Blomqvist, D., Andersson, M. B., Kupper, C., Cuthill, I. C., Kis, J., Lanctot, R. B., Sandercock, B. K., Szekely, T., Wallander, J. & Kempenaers, B. 2002. Genetic similarity between mates and extra-pair parentage in three species of shorebirds. *Nature*, 419, 613–615.
- Bonneaud, C., Chastel, O., Federici, P., Westerdahl, H. & Sorci, G. 2006. Complex Mhc-based mate choice in a wild passerine. Proceedings of the Royal Society of London, Series B, 273, 1111–1116.
- Brooker, M. G., Rowley, I., Adams, M. & Baverstock, P. R. 1990. Promiscuity: an inbreeding avoidance mechanism in a socially monogamous species? *Behavioral Ecology and Sociobiology*, 26, 191–199.
- Brown, J. 1987. Helping and Communal Breeding in Birds. Princeton, New Jersey: Princeton University Press.
- Brown, J. L. 1997. A theory of mate choice based on heterozygosity. Behavioral Ecology, 8, 60–65.
- Carlin, J. B., Wolfe, R., Brown, C. H. & Gelman, A. 2001. A case study on the choice, interpretation and checking of multilevel models for longitudinal binary outcomes. *Biostatistics*, 2, 397–416.
- Carter, C. S., Getz, L. L. & Cohen-Parsons, M. 1986. Relationships between social organisation and behavioural endocrinology in a monogamous mammal. *Advances in the Study of Behavior*, **16**, 109–145.
- **Cohas, A.** 2006. Causes évolutives des paternités hors-couple chez les espèces socialement monogames: l'exemple de la marmotte alpine (*Marmota marmota*). Ph.D. thesis, Université Claude Bernard Lyon 1.
- Cohas, A., Yoccoz, N. G., Da Silva, A., Goossens, B. & Allainé, D. 2006. Extra-pair paternity in the monogamous alpine marmot (*Marmota marmota*): the roles of social setting and female mate choice. *Behavioral Ecology and Sociobiology*, 55, 597–605.
- Cohas, A., Yoccoz, N. G. & Allainé, D. 2007. Extra-pair paternity in alpine marmots, (*Marmota marmota*): genetic quality and genetic diversity effects. *Behavioral Ecology and Sociobiology*, 61, 1081– 1092.
- Cohas, A., Bonenfant, C., Allainé, D. & Gaillard, J.-M. 2007. Are extra-pair young better than within-pair young? A comparison of survival and dominance in alpine marmot. *Journal of Animal Ecology*, **76**, 771–781.
- Coltman, D. W., Pilkington, J. G., Smith, J. A. & Pemberton, J. M. 1999. Parasite-mediated selection against inbred Soay sheep in a free-living island population. *Evolution*, **53**, 1259–1267.
- Creel, S. & Macdonald, D. 1995. Sociality, group-size, and reproductive suppression among carnivores. Advances in the Study of Behavior, 24, 203–257.
- Da Silva, A., Luikart, G., Allainé, D., Gautier, P., Taberlet, P. & Pompanon, F. 2003. Isolation and characterization of microsatellites in European alpine marmots (*Marmota marmota*). *Molecular Ecology Notes*, **3**, 189–190.
- Da Silva, A., Luikart, G., Yoccoz, N. G., Cohas, A. & Allainé, D. 2006. Genetic diversity-fitness correlation revealed by microsatellite analyses in European alpine marmots (*Marmota marmota*). *Conservation Genetics*, 7, 371–382.
- Diggle, P. J., Heagerty, P. J., Liang, K. Y. & Zeger, S. L. 2002. Analysis of Longitudinal Data. Oxford: Oxford University Press.
- Dobson, F. S., Chesser, R. K., Hoogland, J. L., Sugg, D. W. & Foltz,
   D. W. 1997. Do black-tailed prairie dogs minimize inbreeding? *Evolution*, 51, 970–978.
- Eimes, J. A., Parker, P. G., Brown, J. L. & Brown, E. R. 2005. Extrapair fertilization and genetic similarity of social mates in the Mexican jay. *Behavioral Ecology*, 16, 456–460.
- Emlen, S. T. 1982. The evolution of helping: I. An ecological constraints model. *American Naturalist*, **119**, 29–39.
- Foerster, K., Delhey, K., Johnsen, A., Lifjeld, J. T. & Kempenaers,
   B. 2003. Females increase offspring heterozygosity and fitness through extra-pair matings. *Nature*, 425, 714–717.

- Foerster, K., Valcu, M., Johnsen, A. & Kempenaers, B. 2006. A spatial genetic structure and effects of relatedness on mate choice in a wild bird population. *Molecular Ecology*, **15**, 4555–4567.
- Freeman-Gallant, C. R., Meguerdichian, M., Wheelwright, N. T. & Sollecito, S. V. 2003. Social pairing and female mating fidelity predicted by restriction fragment length polymorphism similarity at the major histocompatibility complex in a songbird. *Molecular Ecology*, **12**, 3077–3083.
- Frey-Roos, F. 1998. Geschlechtsspezifisches Abwanderungsmuster beim Alpenmurmeltier (*Marmota marmota*). Ph.D. thesis, Philipps University.
- Goossens, B., Graziani, L., Waits, L. P., Farand, E., Magnolon, S., Coulon, J., Bel, M.-C., Taberlet, P. & Allainé, D. 1998. Extra-pair paternity in the monogamous alpine marmot revealed by nuclear DNA microsatellite analysis. *Behavioral Ecology and Sociobiology*, 43, 281–288.
- Griffith, S. C., Owens, I. P. F. & Thuman, K. A. 2002. Extra pair paternity in birds: a review of interspecific variation and adaptive function. *Molecular Ecology*, **11**, 2195–2212.
- Hackländer, K., Mostl, E. & Arnold, W. 2003. Reproductive suppression in female alpine marmots, *Marmota marmota*. *Animal Behaviour*, 65, 1133–1140.
- Hanslik, S. & Kruckenhauser, L. 2000. Microsatellite loci for two European sciurid species (*Marmota marmota, Spermophilus citellus*). *Molecular Ecology*, 9, 2163–2165.
- Hoffman, J. I., Forcada, J., Trathan, P. N. & Amos, W. 2007. Female fur seals show active choice for males that are heterozygous and unrelated. *Nature*, 445, 912–914.
- Horton, N. J. & Lipsitz, S. R. 1999. Review of software to fit generalized estimating equation regression models. *American Statistician*, 53, 160–169.
- Jennions, M. D. & Macdonald, D. W. 1994. Cooperative breeding in mammals. *Trends in Ecology & Evolution*, 9, 89–93.
- Jombart, T. 2007. The package adegenet for the R software: genetic data handling for multivariate analysis. http://adegenet.r-forge. r-project.org/.
- Kalinowski, S. T., Taper, M. L. & Marshall, T. C. 2007. Revising how the computer program CERVUS accommodates genotyping error increases success in paternity assignment. *Molecular Ecology*, 16, 1099–1106. doi:10.1111/j.1365-294x.2007.03089.x.
- Kempenaers, B. 2007. Mate choice and genetic quality: a review of the heterozygosity theory. Advances in the Study of Behavior, 37, 189–278.
- Kempenaers, B., Congdon, B., Boag, P. T. & Robertson, R. J. 1999. Extrapair paternity and egg hatchability in tree swallows: evidence for the genetic compatibility hypothesis? *Behavioral Ecology*, **10**, 304–311.
- King, W. J. & Allainé, D. 2002. Social, maternal, and environmental influences on reproductive success in female alpine marmots (*Marmota marmota*). Canadian Journal of Zoology, **80**, 2137–2143.
- Klinkicht, M. 1993. Untersuchungen zum Paarungssystem des Alpenmurmeltiers, Marmota M. marmota mittels DNA Fingerprinting. Ph.D. thesis, University of Munich.
- Liang, K. Y. & Zeger, S. L. 1986. Longitudinal data analysis using generalized linear models. *Biometrika*, **73**, 13–22.
- Lynch, M. & Ritland, K. 1999. Estimation of pairwise relatedness with molecular markers. *Genetics*, **152**, 1753–1766.
- Magnolon, S. 1999. Dispersion natale chez la marmotte alpine (*Marmota marmota*). Modalités et effets de quelques facteurs proximaux. Ph.D. thesis, Université de Tours.
- Mitton, J. B. 1993. Theory and data pertinent to the relationship between heterozygosity and fitness. In: *The Natural History of Inbreeding and Outbreeding: Theoretical and Empirical Perspectives* (Ed. by R. Thornhill), pp. 17–41. Chicago: Chicago University Press.

- Perrin, C., Allainé, D. & Le Berre, M. 1993. Socio-spatial organization and activity distribution of the alpine marmot *Marmota marmota*: preliminary results. *Ethology*, **93**, 21–30.
- Pizzari, T. & Birkhead, T. R. 2002. The sexually selected sperm hypothesis: sex-biased inheritance and sexual antagonism. *Biological Reviews*, 77, 183–210.
- Queller, D. C. & Goodnight, K. F. 1989. Estimating relatedness using genetic markers. Evolution, 43, 258–275.
- R Development Core Team. 2007. R: a Language and Environment for Statistical Computing. Vienna: R Foundation for Statistical Computing.
- Rassmann, K., Arnold, W. & Tautz, D. 1994. Low genetic variability in a natural alpine marmot population (*Marmota marmota*, Sciuridae) revealed by DNA fingerprinting. *Molecular Ecology*, **3**, 347– 353.
- Rätti, O., Hovi, M., Lundberg, A., Tegelström, H. & Alatalo, R. V. 1995. Extra-pair paternity and male characteristics in the pied flycatcher. *Behavioral Ecology and Sociobiology*, **37**, 419–425.
- Selander, R. K. 1964. Speciation in the Wrens of the Genus Campylorhyncus. Berkeley: University of California Press.
- Sillero-Zubiri, C., Gottelli, D. & Macdonald, D. W. 1996. Male philopatry, extra-pack copulations and inbreeding avoidance in Ethiopian wolves (*Canis simensis*). *Behavioral Ecology and Sociobiol*ogy, 38, 331–340.
- Stacey, P. & Koenig, W. 1990. Cooperative Breeding in Birds: Long Term Studies of Ecology and Behaviour. Cambridge: Cambridge University Press.
- Tarvin, K. A., Webster, M. S., Tuttle, E. M. & Pruett-Jones, S. 2005. Genetic similarity of social mates predicts the level of extrapair paternity in splendid fairy-wrens. *Animal Behaviour*, 70, 945–955.
- Tregenza, T. F. & Wedell, N. 2000. Genetic compatibility, mate choice and patterns of parentage. *Molecular Ecology*, 9, 1013– 1027.
- Venables, W. N. & Ripley, B. D. 2002. Modern Applied Statistics with S. New York: Springer.
- Westneat, D. F., Sherman, P. W. & Morton, M. L. 1990. The ecology and evolution of extra-pair copulations in birds. *Current Orni*thology, 7, 331–369.
- Wolff, J. O. 1992. Parents suppress reproduction and stimulate dispersal in opposite-sex juvenile white-footed mice. *Nature*, **359**, 409–410.
- Zeger, S. L. & Liang, K. Y. 1986. Longitudinal data analysis for discrete and continuous outcomes. *Biometrics*, 42, 121–130.
- Zeh, J. A. & Zeh, D. W. 1996. The evolution of polyandry: I. Intragenomic conflict and genetic incompatibility. *Proceedings of the Royal Society of London, Series B*, 263, 1711–1717.

## Appendix

To allow polymerase chain reaction multiplexing and subsequent assessment of the allele sizes for each microsatellite locus, we used primer sets labelled with FAM, PET, NED and VIC fluorescent dyes (FAM for SSBibl1, MS45, Ma066 and Ma091, PET for SS-Bibl18, SS-Bibl20, SS-Bibl4, Ma002 and Ma018, NED for SS-Bibl31, MS47, MS53 and Ma002, and VIC for MS41, MS56, MS6 and ST10).

We extracted genomic DNA from 15–30 hairs or from tissue by placing the sample in 50  $\mu$ l lysis buffer (2.0 mM Tris–HCl, 1.5 mM MgCl₂, 25 mM KCl, 0.5% Tween 20, 0.1 mg/ml proteinase K for hairs and 0.4 mg/ml proteinase K for tissue). The samples were incubated at 66°C for 80 min for hairs and at 56°C for 120 min for tissue and then for 20 min at 96°C.

To carry out the polymerase chain reaction (PCR) of the 16 loci we used three 10 µl reactions, Mix1, Mix2 and Mix3, containing 5 µl of Kit PCR (Quiagen, Hilden, Germany) and 1 µl of DNA extract with a DNA concentration of 25-100 ng/ml. In addition, Mix1 contains 0.03 µl of each primer for MS45, 0.1 µl of each primer for SS-Bibl31, MS41and ST10, 0.2 µl of each primer for SS-Bibl18 and SS-Bibl4 and 0.3 µl of each primer of Ma002. Mix2 contains 0.05 µl of each primer for MS56 and MS6, 0.1 µl of each primer for MS53 and Ma091, 0.14 µl of each primer for SSBibl1, and 0.2 ul of each primer for MS47, Ma018 and Ma066. Mix 3 contains 0.2 µl of each primer for SS-Bibl20. Amplifications were carried out in a Mastercycler (Eppendorf, Hamboug, Germany) thermocycler with the following cycling conditions: 15 min at 95°C, then 28 cycles for Mix1 and Mix2 and 35 cycles for Mix3 composed of 30 s denaturing at 94°C, 90 s annealing at 57°C, 60 s extension at 72°C, and finally 30 min at 60°C to ensure complete extension.

We then added 1.5  $\mu$ l of Mix1 and 1.5  $\mu$ l of Mix2 plus 1.5  $\mu$ l Mix3 to 0.15  $\mu$ l of size standard ROX 60-415 and 10  $\mu$ l of formamide and loaded them separately on 5% Long Ranger polyacrylamide gel (Fric). Electrophoresis was run for 3 h on an automated sequencer ABI 3130 (Applied Biosystems Inc., Foster City, CA, U.S.A.) using the size standard, ROX 60-415, to determine allele sizes. Microsatellite patterns were examined with Genemapper 4.0 (Applied Biosystems).